

On the Doubly Truncated Exponential Pareto Distribution

A.A. Akomolafe* and A. Maradesa

¹Department of Statistics, Federal University of Technology Akure, PMB 704, Ondo State, Nigeria.

²The Ibarapa Polytechnic, Eruwa, Oyo State, Nigeria.

E-mail: akomolafe01@yahoo.com*
maradprime1@gmail.com

ABSTRACT

This paper presents a mixture distribution of a new modelling tool which is termed the Doubly Truncated version of exponential Pareto distribution. The distribution allows us to capture some real characteristics of data and it is an important tool for understanding the phenomenon. The various statistical properties of this distribution were fully explored and discussed including the mean, variance, moments, mode, reliability function, and hazard function. The worth of the mixing distribution has been demonstrated by applying it to real life data.

(Keywords: doubly truncated exponential Pareto distribution, reliability function, hazard function, maximum likelihood estimation)

INTRODUCTION

Pareto distribution is widely used for modeling a wide range of data in several fields. It is very worthwhile for analyzing income distribution, actuarial, metrological data and equally preferable for survival analysis. Moreover, it is considered to be the most suitable choice as compared to other parameter distributions in several cases. Kareema, et al (2013) stated that:

$F(x; \alpha, \beta, \theta) = 1 - e^{-\alpha\left(\frac{x}{\beta}\right)^\theta}$ is the cdf of Exponential Pareto Distribution (EPD), then,

$f(x; \alpha, \beta, \theta) = \frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha\left(\frac{x}{\beta}\right)^\theta}$ is the pdf of EPD. from this, we can construct the doubly truncated version of exponential pareto by constraint the values of the random variable X to fall between the interval $a \leq x \leq b$. Where x is the random variable that follows the truncated distribution.

DERIVATION OF THE DOUBLY TRUNCATED EXPONENTIAL PARETO

Let $g(x) = f(x; \alpha, \beta, \theta)$ and $G(x) = F(x; \alpha, \beta, \theta)$, then $f_{DTEPD}(x; \tau) = \frac{g(x)}{G(b)-G(a)}$ (1)

By making necessary substitution, the Pdf of Doubly truncated exponential pareto distribution (DTEPD) will be obtained as (2).= $f_{DTEPD}(x; \alpha, \beta, \theta)$

$$= \frac{\frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha\left(\frac{x}{\beta}\right)^\theta}}{1 - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} - 1 + e^{-\alpha\left(\frac{a}{\beta}\right)^\theta}} = \frac{\frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha\left(\frac{x}{\beta}\right)^\theta}}{e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta}} \quad \text{for } a \leq x \leq b, x > 0, \alpha, \beta, \theta > 0 \quad (2)$$

The (2) above is the probability density function of Doubly-truncated-exponential-pareto distribution. Also, its cdf can be obtained by making necessary substitution in $\frac{G(x)-G(a)}{G(b)-G(a)}$, this can be shown as follows:

$$=F_{DTEPD}(x; \alpha, \beta, \theta) = \frac{1 - e^{-\alpha(\frac{x}{\beta})^\theta} - 1 + e^{-\alpha(\frac{a}{\beta})^\theta}}{1 - e^{-\alpha(\frac{b}{\beta})^\theta} - 1 + e^{-\alpha(\frac{a}{\beta})^\theta}} = \frac{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{x}{\beta})^\theta}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}} \quad (3)$$

RELIABILITY

$$=R(t) = 1 - F_{DTEPD}(x; \alpha, \beta, \theta) = 1 - \left\{ \frac{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{x}{\beta})^\theta}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}} \right\}$$

HAZARD FUNCTION

H(t)=

$$\begin{aligned} \frac{f_{DTEPD}(x; \alpha, \beta, \theta)}{R(t)} &= \frac{\frac{\frac{\alpha\theta}{\beta} (\frac{x}{\beta})^{\theta-1} e^{-\alpha(\frac{x}{\beta})^\theta}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}}}{1 - \left\{ \frac{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{x}{\beta})^\theta}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}} \right\}} = \\ &= \frac{\frac{\alpha\theta}{\beta} (\frac{x}{\beta})^{\theta-1} e^{-\alpha(\frac{x}{\beta})^\theta}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}} \times \frac{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}}{\left\{ \left(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta} \right) \left(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{x}{\beta})^\theta} \right) \right\}} \end{aligned} \quad (4)$$

$$\begin{aligned} &= \frac{\frac{\frac{\alpha\theta}{\beta} (\frac{x}{\beta})^{\theta-1} e^{-\alpha(\frac{x}{\beta})^\theta}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}} - \left\{ \left(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta} \right) \left(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{x}{\beta})^\theta} \right) \right\}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}} \left\{ \frac{\frac{\alpha\theta}{\beta} (\frac{x}{\beta})^{\theta-1} e^{-\alpha(\frac{x}{\beta})^\theta}}{\left(1 - \left(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta} \right) \right)} \right\} \end{aligned} \quad (5)$$

ASYMPTOTIC PROPERTIES

Limit as $x \rightarrow \infty$

$$= \lim_{x \rightarrow \infty} \frac{\frac{\frac{\alpha\theta}{\beta} (\frac{x}{\beta})^{\theta-1} e^{-\alpha(\frac{x}{\beta})^\theta}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}}}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}} = \lim_{x \rightarrow \infty} \frac{\frac{\alpha\theta}{\beta} (\infty)^{\theta-1} \cdot 0}{e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}} = 0$$

Limit as $x \rightarrow 0$

$$= \lim_{x \rightarrow 0} \frac{\frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha\left(\frac{x}{\beta}\right)^\theta}}{e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta}} = \lim_{x \rightarrow 0} \frac{\frac{\alpha\theta}{\beta} (0)^{\theta-1} \cdot 1}{e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta}} = \frac{0}{e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta}} = 0$$

There is an indication of at least one mode, since the limit of $f_{DTEPD}(x; \alpha, \beta, \theta)$ as $x \rightarrow 0$ and $x \rightarrow \infty$ are zero. Then DTEPD has a mode.

RENYI ENTROPY

Proposition 1

The Renyi entropy of DTEPD is given as:

$$\delta_r = \frac{1}{1-r} \log \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-r} \cdot \frac{(\alpha\theta)^r \cdot \beta^{1-r}}{(\alpha r \theta)(\alpha r)^{r-\frac{r}{\theta}+\frac{1}{\theta}-1}} \Gamma\left(\frac{(\theta-1)(r-1)}{\theta} + 1\right)$$

Proof

$$= \delta_r = \frac{1}{1-r} \log \int_0^\infty f^r(x) dx = \frac{1}{1-r} \log \int_0^\infty \left(\frac{\frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha\left(\frac{x}{\beta}\right)^\theta}}{e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta}} \right)^r dx =$$

$$\frac{1}{1-r} \log \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-r} \int_0^\infty \left(\frac{\alpha\theta}{\beta} \right)^r \left(\left(\frac{x}{\beta} \right)^{\theta-1} \right)^r e^{-\alpha r \left(\frac{x}{\beta} \right)^\theta} dx \quad (6)$$

Let $y = \alpha r \left(\frac{x}{\beta} \right)^\theta$, $x = \beta \left(\frac{y}{\alpha r} \right)^{\frac{1}{\theta}}$ and $dx = \frac{dy}{\frac{\alpha r \theta}{\beta} \left(\frac{x}{\beta} \right)^{\theta-1}}$, therefore,

$$= \frac{1}{1-r} \log \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-r} \left(\frac{\alpha\theta}{\beta} \right)^r \int_0^\infty \left(\left(\frac{x}{\beta} \right)^{\theta-1} \right)^r \cdot e^{-y} \frac{dy}{\frac{\alpha r \theta}{\beta} \left(\frac{x}{\beta} \right)^{\theta-1}} \quad (7)$$

$$= \frac{1}{1-r} \log \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-r} \left(\frac{\alpha\theta}{\beta} \right)^r \int_0^\infty \left(\left(\frac{x}{\beta} \right)^{\theta-1} \right)^r \cdot e^{-y} \left(\frac{\alpha r \theta}{\beta} \right)^{-1} \left(\left(\frac{x}{\beta} \right)^{\theta-1} \right)^{-1} dy \quad (8)$$

$$= \frac{1}{1-r} \log \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-r} \left(\frac{\alpha\theta}{\beta} \right)^r \cdot \left(\frac{\alpha r \theta}{\beta} \right)^{-1} \int_0^\infty \left(\left(\frac{x}{\beta} \right)^{\theta-1} \right)^{r-1} \cdot e^{-y} dy \quad (9)$$

$$= \frac{1}{1-r} \log \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-r} \left(\frac{\alpha\theta}{\beta} \right)^r \cdot \left(\frac{\alpha r \theta}{\beta} \right)^{-1} \int_0^\infty \frac{(x^{\theta-1})^{r-1}}{\beta^{(\theta-1)(r-1)}} e^{-y} dy \quad (10)$$

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \cdot \left(\frac{\alpha r\theta}{\beta}\right)^{-1} \beta^{-(r\theta-\theta-r+1)} \int_0^\infty (x^{\theta-1})^{r-1} e^{-y} dy \quad (11)$$

Put for $x = \beta \left(\frac{y}{\alpha r}\right)^{\frac{1}{\theta}}$ in (11)

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \cdot \left(\frac{\alpha r\theta}{\beta}\right)^{-1} \beta^{-(r\theta-\theta-r+1)} \int_0^\infty \left(\left(\beta \left(\frac{y}{\alpha r}\right)^{\frac{1}{\theta}}\right)^{\theta-1}\right)^{r-1} e^{-y} dy \quad (12)$$

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \cdot \left(\frac{\alpha r\theta}{\beta}\right)^{-1} \beta^{-(r\theta-\theta-r+1)} \int_0^\infty \beta^{(r-1)(\theta-1)} \frac{y^{(1-\frac{1}{\theta})(r-1)}}{(\alpha r)^{(1-\frac{1}{\theta})(r-1)}} e^{-y} dy \quad (13)$$

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \cdot \left(\frac{\alpha r\theta}{\beta}\right)^{-1} \beta^{-(r\theta-\theta-r+1)} \frac{\beta^{(r-1)(\theta-1)}}{(\alpha r)^{r-1+\frac{r-1}{\theta}}} \int_0^\infty y^{(1-\frac{1}{\theta})(r-1)} e^{-y} dy \quad (14)$$

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \cdot \left(\frac{\alpha r\theta}{\beta}\right)^{-1} \beta^{-(r\theta-\theta-r+1)} \frac{\beta^{(r-1)(\theta-1)}}{(\alpha r)^{r-1+\frac{r-1}{\theta}}} \int_0^\infty y^{r-\frac{r}{\theta}+\frac{1}{\theta}-1} e^{-y} dy \quad (15)$$

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \cdot \left(\frac{\alpha r\theta}{\beta}\right)^{-1} \beta^{-(r\theta-\theta-r+1)} \frac{\beta^{r\theta-r-\theta+1}}{(\alpha r)^{r-1-\frac{r}{\theta}+\frac{1}{\theta}}} \Gamma\left(\frac{(\theta-1)(r-1)}{\theta} + 1\right) \quad (16)$$

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \left(\frac{\alpha r\theta}{\beta}\right)^{-1} \beta^{-(r\theta+\theta-r-1)} \frac{\beta^{r\theta-r-\theta+1}}{(\alpha r\theta)(\alpha r)^{r-1-\frac{r}{\theta}+\frac{1}{\theta}}} \Gamma\left(\frac{(\theta-1)(r-1)}{\theta} + 1\right) \quad (17)$$

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \left(\frac{\alpha r\theta}{\beta}\right)^{-1} \frac{1}{(\alpha r\theta)(\alpha r)^{r-1-\frac{r}{\theta}+\frac{1}{\theta}}} \Gamma\left(\frac{(\theta-1)(r-1)}{\theta} + 1\right) \quad (18)$$

$$= \frac{1}{1-r} \log(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta})^{-r} \left(\frac{\alpha\theta}{\beta}\right)^r \frac{\beta}{(\alpha r\theta)(\alpha r)^{r-\frac{r}{\theta}+\frac{1}{\theta}-1}} \Gamma\left(\frac{(\theta-1)(r-1)}{\theta} + 1\right) \quad (19)$$

$$= \frac{1}{1-r} \log\left(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}\right)^{-r} (\alpha\theta)^r \cdot \frac{\beta^{-r}\beta}{(\alpha r\theta)(\alpha r)^{r-\frac{r}{\theta}+\frac{1}{\theta}-1}} \Gamma\left(\frac{(\theta-1)(r-1)}{\theta} + 1\right) \quad (20)$$

$$\delta_r = \frac{1}{1-r} \log\left(e^{-\alpha(\frac{a}{\beta})^\theta} - e^{-\alpha(\frac{b}{\beta})^\theta}\right)^{-r} \cdot \frac{(\alpha\theta)^r \cdot \beta^{1-r}}{(\alpha r\theta)(\alpha r)^{r-\frac{r}{\theta}+\frac{1}{\theta}-1}} \Gamma\left(\frac{(\theta-1)(r-1)}{\theta} + 1\right) \quad (21)$$

Equation (21) above represent the Renyi entropy of DTEPD.

β – Entropy

The Renyi entropy of DTEPD is given as:

$$\delta_{\bar{\beta}} = \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \cdot \frac{\beta^{1-\bar{\beta}} (\alpha\theta)^{\bar{\beta}}}{(\alpha\bar{\beta}\theta)(\alpha\bar{\beta})^{\bar{\beta}-\frac{\bar{\beta}}{\theta}+1}} \Gamma\left(\frac{(\theta-1)(\bar{\beta}-1)}{\theta} + 1\right)$$

$$= \delta_{\bar{\beta}} = \frac{1}{\bar{\beta}-1} \int_0^{\infty} f^{\bar{\beta}} dx = \frac{1}{\bar{\beta}-1} \int_0^{\infty} \left(\frac{\frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha \left(\frac{x}{\beta}\right)^{\theta}}}{e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}}} \right)^{\bar{\beta}} dx \quad (22)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \int_0^{\infty} \left(\frac{\alpha\theta}{\beta} \right)^{\bar{\beta}} \left(\left(\frac{x}{\beta}\right)^{\theta-1} \right)^{\bar{\beta}} e^{-\alpha\bar{\beta} \left(\frac{x}{\beta}\right)^{\theta}} dx \quad (23)$$

$$= \text{Let } y = \alpha r \left(\frac{x}{\beta}\right)^{\theta}, x = \beta \left(\frac{y}{\alpha\bar{\beta}}\right)^{\frac{1}{\theta}} \text{ and } dx = \frac{dy}{\frac{\alpha\bar{\beta}\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1}}, \text{ therefore,}$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta} \right)^{\bar{\beta}} \int_0^{\infty} \left(\left(\frac{x}{\beta}\right)^{\theta-1} \right)^{\bar{\beta}} \cdot e^{-y} \frac{dy}{\frac{\alpha\bar{\beta}\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1}} \quad (24)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta} \right)^{\bar{\beta}} \int_0^{\infty} \left(\left(\frac{x}{\beta}\right)^{\theta-1} \right)^{\bar{\beta}} \cdot e^{-y} \left(\frac{\alpha\bar{\beta}\theta}{\beta} \right)^{-1} \left(\left(\frac{x}{\beta}\right)^{\theta-1} \right)^{-1} dy \quad (25)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta} \right)^{\bar{\beta}} \cdot \left(\frac{\alpha\bar{\beta}\theta}{\beta} \right)^{-1} \int_0^{\infty} \left(\left(\frac{x}{\beta}\right)^{\theta-1} \right)^{\bar{\beta}-1} \cdot e^{-y} dy \quad (26)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta} \right)^{\bar{\beta}} \cdot \left(\frac{\alpha\bar{\beta}\theta}{\beta} \right)^{-1} \int_0^{\infty} \frac{(x^{\theta-1})^{\bar{\beta}-1}}{\beta^{(\theta-1)(\bar{\beta}-1)}} e^{-y} dy \quad (27)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta} \right)^{\bar{\beta}} \cdot \left(\frac{\alpha\bar{\beta}\theta}{\beta} \right)^{-1} \beta^{-(\bar{\beta}\theta-\theta-\bar{\beta}+1)} \int_0^{\infty} (x^{\theta-1})^{\bar{\beta}-1} e^{-y} dy \quad (28)$$

$$\text{Put for } x = \beta \left(\frac{y}{\alpha\bar{\beta}}\right)^{\frac{1}{\theta}} \text{ in} \quad (29)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta} \right)^{\bar{\beta}} \cdot \left(\frac{\alpha\bar{\beta}\theta}{\beta} \right)^{-1} \beta^{-(\theta\bar{\beta}-\theta-\bar{\beta}+1)} \int_0^{\infty} \left(\left(\beta \left(\frac{y}{\alpha\bar{\beta}}\right)^{\frac{1}{\theta}} \right)^{\theta-1} \right)^{\bar{\beta}-1} e^{-y} dy \quad (30)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^{\theta}} - e^{-\alpha \left(\frac{b}{\beta}\right)^{\theta}} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta} \right)^{\bar{\beta}} \cdot \left(\frac{\alpha\bar{\beta}\theta}{\beta} \right)^{-1} \beta^{-(\theta\bar{\beta}-\theta-\bar{\beta}+1)} \int_0^{\infty} \beta^{(\bar{\beta}-1)(\theta-1)} \frac{y^{(1-\frac{1}{\theta})(\bar{\beta}-1)}}{(\alpha\bar{\beta})^{(1-\frac{1}{\theta})(\bar{\beta}-1)}} e^{-y} dy \quad (31)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta}\right)^{\bar{\beta}} \cdot \left(\frac{\alpha\bar{\beta}\theta}{\beta}\right)^{-1} \beta^{-(\theta\bar{\beta}-\theta-\bar{\beta}+1)} \frac{\beta^{\theta\bar{\beta}-\bar{\beta}-\theta+1}}{(\alpha r)^{\bar{\beta}-1+\frac{\bar{\beta}}{\theta}+\frac{1}{\theta}}} \int_0^\infty y^{(1-\frac{1}{\theta})(\bar{\beta}-1)} e^{-y} dy \quad (32)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta}\right)^{\bar{\beta}} \cdot \left(\frac{\alpha\bar{\beta}\theta}{\beta}\right)^{-1} \beta^{-(\theta\bar{\beta}-\theta-\bar{\beta}+1)} \frac{\beta^{\theta\bar{\beta}-\bar{\beta}-\theta+1}}{(\alpha r)^{\bar{\beta}-1+\frac{\bar{\beta}}{\theta}+\frac{1}{\theta}}} \int_0^\infty y^{\bar{\beta}-\frac{\bar{\beta}}{\theta}+\frac{1}{\theta}-1} e^{-y} dy \quad (33)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta}\right)^{\bar{\beta}} \cdot \left(\frac{\alpha\bar{\beta}\theta}{\beta}\right)^{-1} \beta^{-\theta\bar{\beta}+\theta+\bar{\beta}-1} \frac{\beta^{\theta\bar{\beta}-\bar{\beta}-\theta+1}}{(\alpha r)^{\bar{\beta}-1+\frac{\bar{\beta}}{\theta}+\frac{1}{\theta}}} \Gamma\left(\frac{(\theta-1)(\bar{\beta}-1)}{\theta} + 1\right) \quad (34)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta}\right)^{\bar{\beta}} \frac{\beta}{(\alpha\bar{\beta}\theta)(\alpha\bar{\beta})^{\bar{\beta}-1+\frac{\bar{\beta}}{\theta}+\frac{1}{\theta}}} \Gamma\left(\frac{(\theta-1)(\bar{\beta}-1)}{\theta} + 1\right) \quad (35)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-\bar{\beta}} \left(\frac{\alpha\theta}{\beta}\right)^{\bar{\beta}} \frac{\beta}{(\alpha\bar{\beta}\theta)(\alpha\bar{\beta})^{\bar{\beta}-1+\frac{\bar{\beta}}{\theta}+\frac{1}{\theta}}} \Gamma\left(\frac{(\theta-1)(\bar{\beta}-1)}{\theta} + 1\right) \quad (36)$$

$$= \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-\bar{\beta}} \cdot \frac{\beta\beta^{-\bar{\beta}}(\alpha\theta)^{\bar{\beta}}}{(\alpha\bar{\beta}\theta)(\alpha\bar{\beta})^{\bar{\beta}-\frac{\bar{\beta}}{\theta}+\frac{1}{\theta}-1}} \Gamma\left(\frac{(\theta-1)(\bar{\beta}-1)}{\theta} + 1\right) \quad (37)$$

$$\delta_{\bar{\beta}} = \frac{1}{\bar{\beta}-1} \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-\bar{\beta}} \cdot \frac{\beta^{1-\bar{\beta}}(\alpha\theta)^{\bar{\beta}}}{(\alpha\bar{\beta}\theta)(\alpha\bar{\beta})^{\bar{\beta}-\frac{\bar{\beta}}{\theta}+\frac{1}{\theta}-1}} \Gamma\left(\frac{(\theta-1)(\bar{\beta}-1)}{\theta} + 1\right) \quad (38)$$

The Equation (38) above is the β – entropy of DTEPD.

MOMENT

$$= E(x^r) = \int_0^\infty x^r f(x) dx = \int_0^\infty x^r \frac{\frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha \left(\frac{x}{\beta}\right)^\theta}}{e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta}} dx \quad (39)$$

$$= \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \int_0^\infty x^r \frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha \left(\frac{x}{\beta}\right)^\theta} dx \quad (40)$$

$$= \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \frac{\alpha\theta}{\beta} \int_0^\infty x^r \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha \left(\frac{x}{\beta}\right)^\theta} dx \quad (41)$$

Substitute for $y = \alpha \left(\frac{x}{\beta}\right)^\theta$; $x = \beta \left(\frac{y}{\alpha}\right)^{\frac{1}{\theta}}$ and $dx = \frac{dy}{\frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1}}$ in

$$\begin{aligned} & \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \frac{\alpha\theta}{\beta} \int_0^\infty x^r \left(\frac{x}{\beta}\right)^{\theta-1} \cdot e^{-y} \frac{dy}{\frac{\alpha\theta(x)}{\beta}^{\theta-1}} = \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - \right. \\ & \left. e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \int_0^\infty x^r e^{-y} dy \end{aligned} \quad (42)$$

$$= \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \int_0^\infty \left(\beta \left(\frac{y}{\alpha}\right)^{\frac{1}{\theta}} \right)^r e^{-y} dy \quad (43)$$

$$= \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \frac{\beta^r}{(\alpha)^{\frac{r}{\theta}}} \int_0^\infty y^{\frac{r}{\theta}} e^{-y} dy = \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \frac{\beta^r}{(\alpha)^{\frac{r}{\theta}}} \Gamma\left(\frac{r}{\theta} + 1\right) \quad (44)$$

$$E(x^r) = \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \left(\frac{\beta}{\theta\sqrt{\alpha}} \right)^r \Gamma\left(\frac{r}{\theta} + 1\right), r = 1, 2, 3 \dots \quad (45)$$

$$= E(x) = \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \frac{\beta}{\theta\sqrt{\alpha}} \Gamma\left(\frac{1}{\theta} + 1\right) \quad (46)$$

$$= E(x^2) = \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \left(\frac{\beta}{\theta\sqrt{\alpha}} \right)^2 \Gamma\left(\frac{2}{\theta} + 1\right) \quad (47)$$

$$E(x^3) = \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \left(\frac{\beta}{\theta\sqrt{\alpha}} \right)^3 \Gamma\left(\frac{3}{\theta} + 1\right) \quad (48)$$

$$= E(x^4) = \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \left(\frac{\beta}{\theta\sqrt{\alpha}} \right)^4 \Gamma\left(\frac{4}{\theta} + 1\right) \quad (49)$$

$$\begin{aligned} \text{Var}(x) &= E(x^2) - (E(x))^2 = \\ & \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \left(\frac{\beta}{\theta\sqrt{\alpha}} \right)^2 \Gamma\left(\frac{2}{\theta} + 1\right) - \left(\left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \frac{\beta}{\theta\sqrt{\alpha}} \Gamma\left(\frac{1}{\theta} + 1\right) \right)^2 \end{aligned} \quad (50)$$

MAXIMUM LIKELIHOOD ESTIMATION

$$= Lf(x; \tau) = \prod_{i=1}^n \frac{\frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha\left(\frac{x}{\beta}\right)^\theta}}{e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta}} = \prod_{i=1}^n \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-1} \frac{\alpha\theta}{\beta} \left(\frac{x}{\beta}\right)^{\theta-1} e^{-\alpha\left(\frac{x}{\beta}\right)^\theta} \quad (51)$$

$$= Lf(x; \tau) = \left(e^{-\alpha\left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha\left(\frac{b}{\beta}\right)^\theta} \right)^{-n} \alpha^n \theta^n \beta^{-n} \left(\frac{\prod_{i=1}^n x_i}{\beta^n} \right)^{\theta-1} e^{-\alpha \frac{\sum_{i=1}^n x_i}{\beta^\theta}} \quad (52)$$

=ln

$$L_f(\mathbf{x}; \boldsymbol{\tau}) = -n \ln \left(e^{-\alpha \left(\frac{a}{\beta}\right)^\theta} - e^{-\alpha \left(\frac{b}{\beta}\right)^\theta} \right) + n \ln \alpha + n \ln \theta - n \ln \beta + (\theta - 1) \ln \prod_{i=1}^n x_i - n(\theta - 1) \ln \beta - \alpha \frac{\sum_{i=1}^n x_i^\theta}{\beta^\theta}$$

(53)

From (53), the parameters of DTEPD can be estimated via numerical method (Newton Raphson algorithm) and the information matrix can be used to test for the stability of the parameters and thus contributing to the reliability of the model over its parent distribution in modeling pattern or behavior of certain random process or stochastic phenomenon at a specified interval ($a \leq x \leq b$).

CONCLUSIONS

By compounding two or more probability distributions, we get the corresponding hybrid distribution with increased number of parameters which is believed to give the newly compounded distribution more flexibility, consistency, stability, sufficiency uniqueness and wider applicability as compare to its parent distribution. Therefore, the hybrid Doubly Truncated exponential Pareto distribution is said to have vast applicability extending beyond modeling simple Brownian movement but can be used to model statistical behavior of stochastic processes such as studying the consumers' buying behavior and in complex epidemiological studies because of its increased number of parameters which give the hybrid distribution more flexibility to model many stochastic phenomena.

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